

Handout¹ 8

Nov 7, 2022

[Topics]:

Limiting Theorems

1 Some Inequalities**Markov Inequality**

Suppose a random variable X has finite k 'th moment, i.e. $E(|X|^k) < \infty$. Then for any $\epsilon > 0$,

$$P(|X| \geq \epsilon) \leq \frac{E(|X|^k)}{\epsilon^k}$$

Proof

Assume X is a continuous random variable with pdf $f(x)$, then

$$\begin{aligned} P(|X| \geq \epsilon) &= \int_{|X| \geq \epsilon} f(x) dx \\ &\leq \int_{|X| \geq \epsilon} \left(\frac{|x|}{\epsilon}\right)^k f(x) dx, \quad \text{as } |x| \geq \epsilon \\ &= \frac{1}{\epsilon^k} \int_{|X| \geq \epsilon} |x|^k f(x) dx \\ &\leq \frac{1}{\epsilon^k} \int_{-\infty}^{+\infty} |x|^k f(x) dx \\ &= \frac{1}{\epsilon^k} \times E(|X|^k) \\ &= \frac{E(|X|^k)}{\epsilon^k} \end{aligned}$$

The proof for discrete random variable case is similar.

Corollary of Markov Inequality

Suppose a random variable X satisfies the condition that $E(X^2) < \infty$, then for any $a > 0$, we have:

$$P(|X| \geq a) \leq \frac{E(X^2)}{a^2}$$

Proof For Markov inequality, take $k = 2$.

Chebyshev Inequality

Suppose a random variable X satisfies the condition that $E(X^2) < \infty$, then for any $\epsilon > 0$, we have:

$$P(|X - E(X)| \geq \epsilon) \leq \frac{\text{Var}(X)}{\epsilon^2}$$

Proof Let $Y = X - E(X)$ and apply the above inequality:

$$\begin{aligned} P(|X - E(X)| \geq \epsilon) &\leq \frac{E(X - E(X))^2}{\epsilon^2} \\ &= \frac{\text{Var}(X)}{\epsilon^2} \end{aligned}$$

¹This handout is made by Hongkai Wang for Econ 329 Economic Statistics. It is adapted from Professor Wiseman's lectures, however, all errors are mine.

2 Weak Law of Large Numbers

Suppose X_1, X_2, \dots, X_n are independent random variables with common mean μ and common finite variance σ^2 . Let $\bar{X}_n = \frac{1}{n} \sum_{i=1}^n X_i$ denotes the sample mean. Then for any $\epsilon > 0$, we have

$$\lim_{n \rightarrow \infty} P(|\bar{X}_n - \mu| > \epsilon) = 0$$

Proof

$$\begin{aligned} E(\bar{X}_n) &= E\left(\frac{1}{n} \sum_{i=1}^n X_i\right) \\ &= \frac{1}{n} \sum_{i=1}^n E(X_i) \\ &= \mu \end{aligned}$$

And

$$\begin{aligned} \text{Var}(\bar{X}_n) &= \text{Var}\left(\frac{1}{n} \sum_{i=1}^n X_i\right) \\ &= \left(\frac{1}{n}\right)^2 \sum_{i=1}^n \text{Var}(X_i) \\ &= \frac{\sigma^2}{n} \end{aligned}$$

Now apply Chebyshev's inequality,

$$\begin{aligned} P(|\bar{X}_n - E(\bar{X}_n)| \geq \epsilon) &\leq \frac{\text{Var}(\bar{X}_n)}{\epsilon^2} \\ P(|\bar{X}_n - \mu| \geq \epsilon) &\leq \frac{\sigma^2}{n\epsilon^2} \end{aligned}$$

As $n \rightarrow \infty$, we have $\lim_{n \rightarrow \infty} P(|\bar{X}_n - \mu| > \epsilon) = 0$ for any $\epsilon > 0$.

Interpretation

For any observations of a distribution, as the size of the observation sample goes larger, the sample mean will converge to the mean of the distribution.

3 Central Limit Theorem

Suppose X_1, X_2, \dots, X_n are independent random variables with common mean μ and common finite variance σ^2 . Let $\bar{X}_n = \frac{1}{n} \sum_{i=1}^n X_i$ denotes the sample mean. Then for each fixed x , we have

$$\lim_{n \rightarrow \infty} P\left(\frac{\sqrt{n}(\bar{X}_n - \mu)}{\sigma} \leq x\right) = \Phi(x)$$

where $\Phi(x)$ denotes the cdf of standard normal distribution.

Interpretation

For any observations of a distribution, as the size of the observation sample goes larger, the sample mean will behave as a normal distribution.

As in section 2, $E(\bar{X}_n) = \mu$ and $\text{Var}(\bar{X}_n) = \frac{\sigma^2}{n}$. Then

$$\begin{aligned} E\left(\frac{\sqrt{n}(\bar{X}_n - \mu)}{\sigma}\right) &= \frac{\sqrt{n}}{\sigma} E(\bar{X}_n - \mu) \\ &= \frac{\sqrt{n}}{\sigma} \times 0 \\ &= 0 \end{aligned}$$

And

$$\begin{aligned} \text{Var}\left(\frac{\sqrt{n}(\bar{X}_n - \mu)}{\sigma}\right) &= \left(\frac{\sqrt{n}}{\sigma}\right)^2 \text{Var}(\bar{X}_n - \mu) \\ &= \frac{n}{\sigma^2} \times \frac{\sigma^2}{n} \\ &= 1 \end{aligned}$$